## NOTES

## REPRODUCTIVE PARAMETERS OF THE OFFSHORE SPOTTED DOLPHIN, A GEOGRAPHICAL FORM OF stenella attenuata, IN THE EASTERN TROPICAL PACIFIC, 1973-75

Perrin et al. (1976) presented estimates of reproductive parameters of the offshore population of Stenella attenuata in the eastern Pacific based on data collected in 1968-73, inclusive. The sample included 3,527 specimens. Only the 1973 sample $(2,036)$ was putatively cross-sectional with respect to age and sex structures of the kill; in earlier years, adult females were selected for examination. The purpose of this paper is to present analyses of samples collected in uniform fashion in 1973, 1974, and 1975, updating the prior report and providing a uniformly developed, albeit short, time series of annual estimates.

Methods and Materials
The data and specimens were collected by NMFS biological technicians aboard commercial tuna vessels. Data collection procedures were the same as described by Perrin et al. (1976). Data on S. attenuata were collected on 24 cruises in 1973, 33 in 1974, and 32 in 1975.

The total sample includes 6,243 specimens, 6,168 from precisely known localities (Figure 1). Because of the seasonal nature of the fishery, the sample is heavily biased toward the first half of the calendar year with practically no coverage of the summer months (Table 1).

Laboratory procedures were the same as reported by Perrin et al. (1976), but the analytical methods differed slightly. In calculating gross annual reproductive rate (proportion female $\times$ proportion of total females which are reproductive


FIGURE 1.-Sample of the offshore spotted dolphin, Stenella attenuata, collected in 1973-75, by $5^{\circ}$ squares.

TABLE 1.-The sample of the offshore spotted dolphin, Stenella attenuata, by sex, year, and month, 1973-75. Date of capture was not available for 18 of the total 6,243 specimens collected.

| Month | 1973 |  | 1974 |  | 1975 |  | Total |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | 6 | 9 | \% | 9 | $\delta$ | 9 | $\delta$ | 9 |
| Jan. | 267 | 326 | 239 | 300 | 395 | 442 | 901 | 1,068 |
| Feb. | 200 | 231 | 428 | 532 | 249 | 312 | 877 | 1,075 |
| Mar. | 137 | 210 | 66 | 72 | 133 | 153 | 336 | 435 |
| Apr. | 41 | 46 | 35 | 42 | 135 | 183 | 211 | 271 |
| May | 85 | 156 | 5 | 2 | 34 | 35 | 124 | 193 |
| June | 56 | 69 | 36 | 9 | 30 | 47 | 122 | 125 |
| July | 0 | 0 | 0 | 0 | 6 | 12 | 6 | 12 |
| Aug. | 0 | 0 | 11 | 13 | 2 | 4 | 13 | 17 |
| Sept. | 0 | 0 | 1 | 4 | 0 | 0 | 1 | 4 |
| Oct. | 5 | 16 | 0 | 0 | 31 | 51 | 36 | 67 |
| Nov. | 72 | 103 | 20 | 48 | 32 | 39 | 124 | 190 |
| Dec. | 8 | 9 | 0 | 0 | 0 | 0 | 8 | 9 |
| Totals |  | $\overline{1,166}$ | 841 | $\begin{array}{r} \overline{1,022} \\ , 863 \\ \hline \end{array}$ | 1.047 | $\begin{array}{r} 1,278 \\ 2,325 \end{array}$ | 2,759 | $\begin{array}{r} 3,466 \\ 6,225 \end{array}$ |

$\times$ annual pregnancy rate), Perrin et al. (1976) estimated the proportion of adult females which were reproductive from coloration phase data, based on a subsample of data on percentage mature in the various coloration phases ("mottled" and "fused-adult"). In the present study, a much larger sample of complete reproductive data was available; therefore, the proportion of total females which were reproductive was estimated directly from that sample. Specimens for which ovarian data were lacking or incomplete were allocated to mature or immature categories based on a length criterion. Average length at attainment of sexual maturity was estimated as that length ( 177 cm ) at which the number of shorter but mature specimens in the sample equals the number of longer but immature specimens.

## Results and Discussion

## Calving Cycle and Pregnancy Rate

The calving cycle, for purposes of analyzing field data, can be divided into three phases: 1) pregnancy, 2) lactation, and 3) "resting" (a catch-
all phase for animals neither pregnant nor lactating, which includes females truly resting, i.e., not ovulating because of being between cycles, those which have just ovulated but have not become pregnant, some with extremely small embryos missed in dissection, those which recently aborted, and those which have prematurely terminated lactation due to death of the suckling calf).

We estimated the length of the cycle (and pregnancy rate) in two ways: 1) based on the reproductive structure of the sample of adult females, i.e., based on the assumption that the samples are not biased with respect to reproductive phase, and that the proportion of a sample of mature females in a particular phase is directly proportional to the relative length of that phase, using the previously estimated (Perrin et al. 1976) length of gestation ( 11.5 mo ) as a time calibration, and 2) based on the estimate of length of gestation and a largely independent estimate of length of lactation.
The first estimate was based on data for 1,876 females classified as pregnant, lactating, pregnant and lactating, "resting," or postreproductive (Table 2). The "resting" females were further subdivided into those with and without a corpus luteum. As discussed above, some proportion of those with a corpus luteum can be assumed to represent females not truly resting (with a corpus luteum of infertile ovulation). In the total sample of 3,443 females, 61 were simultaneously pregnant and lactating ( $6.1 \%$ of the lactating females). Minor differences between the numbers in Table 2 and in table 8 of Perrin et al. (1976) reflect increase of the 1973 sample by eight specimens and reexamination and reevaluation of the materials.
Subtraction of the postreproductive females from the aggregate of mature females of determined reproductive condition and allocation of the females both pregnant and lactating to both

TABLE 2.-Reproductive condition of 3,469 female offshore spotted dolphins, Stenella attenuata, collected 1973-75.

| Condition | 1973 |  | 1974 |  | 1975 |  | 1973-75 pooled |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | No. | \% | No. | \% | No. | \% | No. | \% |
| Sexually immature | 522 | 45.0 | 465 | 45.9 | 580 | 45.7 | 1,567 | 45.5 |
| Sexually mature: |  |  |  |  |  |  |  |  |
| Condition undetermined | 58 | 5.0 | 60 | 5.9 | 191 | 15.0 | 309 | 9.0 |
| Pregnant only | 232 | 20.0 | 122 | 12.1 | 119 | 9.4 | 473 | 13.7 |
| Pregnant and lactating | 16 | 1.4 | 23 | 2.3 | 22 | 1.7 | 61 | 1.8 |
| Lactating only | 22.6 | 19.5 | 256 | 25.3 | 264 | 20.8 | 746 | 21.7 |
| "Resting" |  |  |  |  |  |  |  |  |
| With corpus luteum | 34 | 2.9 | 32 | 3.2 | 28 | 2.2 | 94 | 2.7 |
| Without corpus luteum | 66 | 5.7 | 48 | 4.7 | 64 | 5.0 | 178 | 5.2 |
| Postreproductive | 7 | 0.6 | 6 | 0.6 | 2 | 0.2 | 15 | 0.4 |
| Totals | 1,161 | 100.0 | 1,012 | 100.0 | 1,270 | 100.0 | 3,443 | 100.0 |

categories provides estimates of the proportions of reproductive females in the three phases of the cycle and, comparing the proportions, of the relative lengths of the phases. Estimated average length of the phases and the total cycle can then be calculated for each $1-\mathrm{yr}$ sample and for the pooled samples, based on the relative lengths of the phases and on the estimated gestation period of 11.5 mo (Method 1 in Table 3). The estimates of average length of cycle thus derived trend from 27.3 mo in 1973 to 42.3 mo in 1975, due to increase in the estimated length of lactation from 11.2 mo to 23.3 mo .

Annual pregnancy rate under Method 1 (also in Table 3) is calculated as proportion of reproductive females pregnant divided by the length of gestation ( 0.958 yr ). The reciprocal of annual pregnancy rate is the estimate of average calving interval.

In the second method of calculating length of calving cycle, we estimated length of lactation by assuming that a suckling calf existed in the samples for each lactating female. Under this assumption, the length at which the cumulative frequency of calves in a sample equals the number of lactating females should be the average length at weaning (from which, using the length-age equations published by Perrin et al. (1976), the average age at weaning can be calculated). If the length of lactation increases, the average length

Table 3.-Estimates of lengths of reproductive phases, pregnancy rate, and calving interval under two methods of estimating length of calving cycle (see text) of the offshore spotted dolphin, Stenella attenuata, 1973-75.

| Item | 1973 | 1974 | 1975 | 1973-75 pooled |
| :---: | :---: | :---: | :---: | :---: |
| Sample size (no.) | 574 | 481 | 497 | 1,552 |
| Pregnancy (mo) | 11.5 | 11.5 | 11.5 | 11.5 |
| Lactation (mo): |  |  |  |  |
| Method 1 | 11.2 | 21.9 | 23.3 | 17.4 |
| Method 2 (Hyp. II) | 11.2 | 12.4 | 12.1 | 11.9 |
| "Resting" (Method 1 ) | 4.6 | 6.4 | 7.5 | 5.9 |
| Sum of phases: |  |  |  |  |
| Months | 27.3 | 39.8 | 42.3 | 34.8 |
| Years | 2.28 | 3.32 | 3.53 | 2.90 |
| Method 2 |  |  |  |  |
| Months | 27.3 | 28.5 | 28.2 | 28.0 |
| Years | 2.28 | 2.38 | 2.35 | 2.33 |
| Annual pregnancy rate (APR): |  |  |  |  |
| Method 1 | 0.452 | 0.314 | 0.296 | 0.359 |
| Method 2 | 0.472 | 0.459 | 0.461 | 0.463 |
| Calving interval (1/APR): |  |  |  |  |
| Method 1 |  |  |  |  |
| Years | 2.21 | 3.18 | 3.38 | 2.79 |
| Months | 26.5 | 38.2 | 40.5 | 33.4 |
| Method 2 |  |  |  |  |
| Years | 2.12 | 2.18 | 2.17 | 2.16 |
| Months | 25.4 | 26.2 | 26.0 | 25.9 |

at weaning estimated by this method should increase concomitantly. The calculated length at weaning did not increase sharply between years (Table 4). Under Hypothesis II of Perrin et al. (1976) of the rate of deposition of dentinal growth layers (two in first year and one per year there-after-the most likely alternative), the estimated length of lactation ranges from 11.2 mo in 1973 to 12.4 mo in 1974 . To arrive at estimates of the total length of the calving cycle under Method 2, we used the estimate of time spent in the "resting" phase under Method 1 for 1973 (the year for which the two estimates of length of lactation coincide exactly) or 4.6 mo , for each of the three annual estimates. This estimate is based on the assumptions under Method 1 but must suffice as a first approximation. In estimating pregnancy rate (as reciprocal of calving interval)-Table 3-overlapping cycles were taken into consideration by adjusting the effective length of lactation downward by a factor equal to the percentage of lactating females also pregnant.

The 1973 estimate of length of lactation (and length of cycle, pregnancy rate, and calving interval) is very close to that obtained by Method 1 above ( 11.2 mo ), but the two sets of estimates diverge sharply thereafter. The first method could be invalid and cause diverging estimates if 1) lactating females (and their nursing calves) were overrepresented in the samples for 1974 and 1975 or, conversely, 2) either (or both) pregnant or "resting" females were underrepresented. The first situation could obtain if lactating females and their accompanying calves are more likely to be captured and killed in the net because of limited endurance and ability to escape of the calf, certainly less than those of adults, and the strength of the mother-calf bond. The second method could yield erroneous estimates if 1) nursing calves were overrepresented in the samples or,

TABLE 4.-Estimates of length of lactation in the offshore spotted dolphin, Stenella attenuata, based on the cumulative calf length/lactating females method (see text) 1973-75.

| Sample | Lactating females ${ }^{1}$ (no.) | Length ${ }^{2}$ (cm) | Length of lactation |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  |  | Under hypothesis |  |  |
|  |  |  | Growth layers | $\begin{gathered} \overline{1} \\ \text { (mo) } \end{gathered}$ | $\begin{gathered} 11 \\ (\mathrm{mo}) \end{gathered}$ | $\overline{\mathrm{IIII}}$ |
| 1973 | 259 | 135.8 | 1.86 | 11.2 | 11.2 | 11.2 |
| 1974 | 301 | 138.5 | 2.03 | 12.2 | 12.4 | 12.2 |
| 1975 | 376 | 138.2 | 2.01 | 12.1 | 12.1 | 12.1 |
| $\begin{array}{r} \text { 1973-75 } \\ \text { pooled } \end{array}$ | 936 | 137.8 | 1.98 | 11.9 | 11.9 | 11.9 |

Includes mature females ( $\geqslant 177 \mathrm{~cm}$ ) without lactation data prorated to lactating and nonlactating based on proportions in sample with lactation data.
${ }^{2}$ Length at which cumulative number of calves $=$ number of lactating females.
conversely, 2) lactating females were underrepresented. Ongoing analyses of data for the spotted dolphin, S. attenuata (J. E. Powers pers. commun.), indicate that small calves are probably overrepresented in small single-set samples. In addition, the absence of sharp change in length of calves at weaning as estimated by the cumulative-calves method speaks against the alternative explanation of development between the years of differential bias against calves and lactating females. The balance of evidence favors the first alternative above, that of progressive overrepresentation of both nursing calves and lactating females as the average number of animals encircled has increased and the average number killed per net haul has decreased ${ }^{1}$ accentuating the factor of differential stamina.

## Gross Annual Reproduction

Estimates of gross annual reproductive rates can be calculated based on the two methods of estimating pregnancy rate (Table 5). It must be noted that if, as discussed above, small calves are overrepresented in small samples (which make up most of the aggregate sample), then pregnancy rate (and, therefore, gross annual reproductive rate) under Method 1 are underestimated to an unknown, but probably small, degree. This factor, of course, would also cause overestimation of the proportion of the total sample female and the pro-

[^0]portion of total females which are reproductive, causing a countering overestimation of gross annual reproduction of unknown, but again probably small, size.

Standard errors are attached to the various estimates where sample size $\geqslant 100$, under the assumption that the binomial distribution tends to normality in large samples (Bailey 1959), allowing calculation of the standard error as:

$$
\mathrm{SE}=\sqrt{p(1-p) / n}
$$

where $p=$ proportion (estimate of parameter)
$n=$ sample size .
Although gross annual reproductive rate as calculated in Table 5 is a product of three estimates, it can be calculated directly from the total sample (number of females pregnant $\div$ total number of males and females), to yield the same estimate and allowing estimation of the variance by the above method. The total sample size was adjusted downward by a factor equal to the proportion of mature females in unknown reproductive condition. The effect on the variance by the constant used to adjust the pregnancy rate to an annual rate was ignored, because the constant ( 11.5 mo gestation $\div 12 \mathrm{mo}$, or 0.958 ) is close to unity.

The estimates of pregnancy rate (and gross annual reproductive rate) for 1973 and 1974 based on structure of the samples (Method 1) are significantly different from each other (using $\pm 2 \mathrm{SE}$ as an approximation of a $95 \%$ confidence interval), and the estimate for 1975, although not statistically different from that for 1974, continues the trend. The estimates based on independent esti-

TABLE 5.-Calculation of estimates of gross annual reproductive rate of offshore spotted dolphin, Stenella attenuata, for 1973-75, using two alternative estimates of pregnancy rate (see text). Standard error follows estimate (see text).

| Year | A | B |  |  | A $\times$ | $\times \mathrm{C}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Proportionfemale | Proportion of females reproductive | Annual pregnancy rate |  | Gross annual reproductive rate |  |
|  |  |  | Method 1 | Method 2 | Method 1 | Method 2 |
| 1971 |  |  | 0.546 |  |  |  |
|  |  |  | (86) |  |  |  |
| 1972 |  |  | $\begin{gathered} 0.465 \pm 0.023 \\ (455) \end{gathered}$ |  |  |  |
| 1973 | $\begin{gathered} 0.572 \pm 0.011 \\ (2,037) \end{gathered}$ | $\begin{gathered} 0.544 \pm 0.015 \\ (1,161) \end{gathered}$ | $\begin{gathered} 0.452 \pm 0.021 \\ (574) \end{gathered}$ | $\begin{gathered} 0.472 \pm 0.021 \\ (574) \end{gathered}$ | $\begin{gathered} 0.141 \pm 0.008 \\ (2,036) \end{gathered}$ | $\begin{gathered} 0.147 \pm 0.008 \\ (1,934) \end{gathered}$ |
| 1974 | $\begin{gathered} 0.548 \pm 0.012 \\ (1,863) \end{gathered}$ | $\begin{gathered} 0.535 \pm 0.016 \\ (1,012) \end{gathered}$ | $\begin{gathered} 0.314 \pm 0.021 \\ (481) \end{gathered}$ | $\begin{gathered} 0.459 \pm 0.023 \\ (481) \end{gathered}$ | $\begin{gathered} 0.092 \pm 0.007 \\ (1,860) \end{gathered}$ | $\begin{gathered} 0.135 \pm 0.008 \\ (1,750) \end{gathered}$ |
| 1975 | $\begin{gathered} 0.559 \pm 0.010 \\ (2,321) \end{gathered}$ | $\begin{gathered} 0.542 \pm 0.014 \\ (1,270) \end{gathered}$ | $\begin{gathered} 0.296 \pm 0.020 \\ (523) \end{gathered}$ | $\begin{gathered} 0.461 \pm 0.022 \\ (523) \end{gathered}$ | $\begin{gathered} 0.087 \pm 0.006 \\ (2,321) \end{gathered}$ | $\begin{gathered} 0.140 \pm 0.007 \\ (2,001) \end{gathered}$ |
| 1973-75 | $\begin{gathered} 0.560 \pm 0.006 \\ (6,221) \\ \hline \end{gathered}$ | $\begin{gathered} 0.541 \pm 0.008 \\ (3,443) \\ \hline \end{gathered}$ | $\begin{gathered} 0.359 \pm 0.012 \\ (1,578) \end{gathered}$ | $\begin{gathered} 0.463 \pm 0.012 \\ (1,578) \\ \hline \end{gathered}$ | $\begin{gathered} 0.109 \pm 0.004 \\ (6,243) \end{gathered}$ | $\begin{gathered} 0.140 \pm 0.004 \\ (5,685) \end{gathered}$ |

mates of lengths of gestation and lactation also trend downward, but the year-to-year differences are not statistically significant. As discussed above, the balance of evidence favors the (relatively nonvarying) estimates based on Method 2.

Although adequate data for estimating sex ratio and proportion of total females which were reproductive in 1971 and 1972, using the methods employed here, or for estimating pregnancy rate using Method 2, are not available because of selection of adult females for dissection, the estimates of pregnancy rate (using Method 1) for those two years are included in Table 5. The sample for 1971 is too small to allow direct statistical comparison with the estimates for later years, but the 1972 estimate is not significantly different from the estimates for 1973, reinforcing the suggestion that a major shift in population structure or (more likely) in degree of representativeness of the kill or the sample occurred in 1974.

In summary, the balance of evidence indicates that, in management of the dolphin/tuna situation, changes in the structure of the dolphin kill, per se, should not be taken to necessarily mean that parallel changes in reproductive rates have occurred. The changes more likely represent sampling artifacts caused by changes in the fishing and dolphin rescue operations.

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THE UPTAKE, DISTRIBUTION, AND DEPURATION OF ${ }^{14} \mathrm{C}$ BENZENE AND ${ }^{14} \mathrm{C}$ TOLUENE IN PACIFIC HERRING, clupea harengus pallasi

This note is a sequel to Korn et al. (1976), where uptake, distribution, and depuration of ${ }^{14} \mathrm{C}$ benzene were examined in striped bass, Morone saxatilis, and northern anchovy, Engraulis mor$d a x$. Like benzene, toluene is a prevalent, watersoluble, and toxic monoaromatic component of petroleum and associated products. According to Anderson et al. (1974a), toluene is second only to benzene as the most abundant aromatic oil component in the water-soluble extracts of southern Louisiana and Kuwait crude oils (6.75-3.36 $\mu \mathrm{l} /$ liter benzine; 4.13-3.62 $\mu \mathrm{l}$ /liter toluene, respectively).

Although levels of the volatile aromatics are thought to be low in areas subject to chronic oil exposure, few actual measurements have been made. Further, if fish can accumulate benzene and if energy is required to metabolize, detoxify, and depurate these aromatics, long-term physiological and population effects are possible.

In this study, a comparison of the uptake, distribution, and depuration of ${ }^{14} \mathrm{C}$ benzene and ${ }^{14} \mathrm{C}$ toluene, at a low sublethal concentration [ 100 parts per billion (ppb)], was undertaken to determine which of these prevalent aromatics may pose the greatest problem. It was hypothesized that, although toluene is less soluble in seawater (Anderson et al. 1974a), it may be more toxic and exhibit greater accumulation levels and persistence. Our previous work with striped bass and northern anchovy indicated other tissues that should be examined, such as kidney, pyloric caeca, gonad, and intestine, and in the present comparison, residues in the additional tissues were measured. Pacific herring, Clupea harengus pallasi, were selected as test animals because of their importance as estuarine and nearshore forage fish for many important recreational and commercial species, including striped bass and chinook salmon.

## Methods

Pacific herring were obtained from a San Francisco Bay bait dealer and were transported directly to the Tiburon Laboratory dock. The fish were acclimated under test conditions for at least 2


[^0]:    ${ }^{1}$ Staff, Porpoise/Tuna Interaction Program, Oceanic Fisheries Resources Division. 1975. Progress of research on porpoise mortality incidental to tuna purse-seine fishing for fiscal year 1975. SWFC Admin. Rep., Natl. Mar. Fish. Serv., La Jolla, Calif., LJ-75-68, 98 p. (Unpubl. rep.)

